



Posterior Inference in Latent Gaussian Models Using Manifold MCMC Methods

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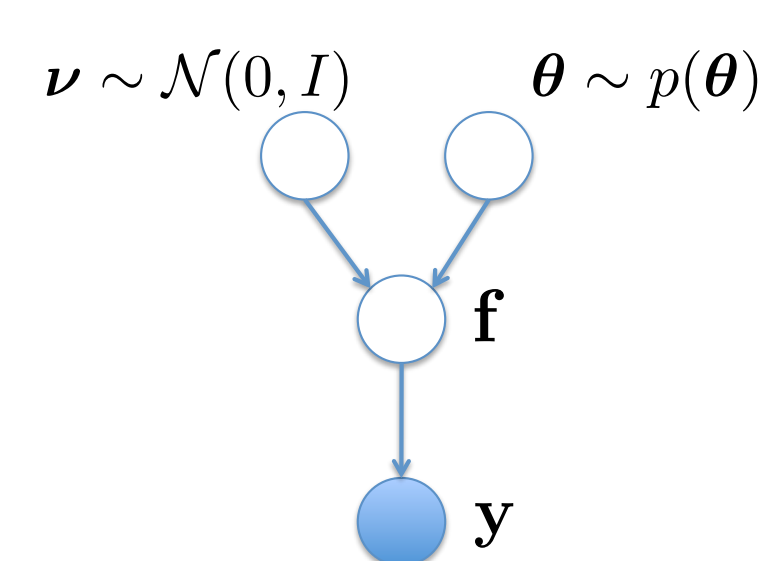
Scope of this work

In this work, we study the inference problem in latent Gaussian models. Efficiently sampling from the posterior distribution of the latent process and hyperparameters is complex because of their strong coupling. We consider a set of recently proposed MCMC methods based on the natural geometry of the underlying statistical model to achieve efficient sampling.

Latent Gaussian Models - LGMs

- ▶ Let $X = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$ be a set of n covariates $\mathbf{x}_i \in \mathbb{R}^d$, associated with observed responses $\mathbf{y} = y_1, \dots, y_n$.
- ▶ Let k be the covariance function parameterized by hyperparameters θ
- ▶ Consider the following general form of latent Gaussian models to model the generative process of the observed \mathbf{y} (given X).

$p(\theta)$	prior θ
$K = LL^T$	covariance matrix
$p(\nu) \sim \mathcal{N}(0, I)$	whitened latent transformation
$\mathbf{f} = L\nu$	
\downarrow	
$p(\mathbf{f} \theta) = \mathcal{N}(\mathbf{f} \theta, K)$	prior latent \mathbf{f}
$p(\mathbf{y} \mathbf{f}) = \mathcal{E}(\mathbf{y} \zeta(\mathbf{f}))$	likelihood



- ▶ Distribution of the observed random variables \mathbf{y} : exponential family \mathcal{E} with natural parameters given by a transformation of the latent variables $\zeta(\mathbf{f})$.

$$\mathcal{E}(\mathbf{y}|\zeta(\mathbf{f})) = h(\mathbf{y})g(\zeta(\mathbf{f})) \exp(\zeta(\mathbf{f})^T \mathbf{u}(\mathbf{y}))$$

- ▶ The inference problem amounts in obtaining the posterior distribution over the parameters and use it to compute the predictive distribution

$$p(y_*|\mathbf{y}) = \int \int \int p(y_*|f_*)p(f_*|\mathbf{f}, \theta)p(\mathbf{f}, \theta|\mathbf{y})df_*d\mathbf{f}d\theta$$

which is intractable. In this work we apply Markov Chains Monte Carlo (MCMC) methods to integrate out \mathbf{f} and θ .

Manifold MCMC

- ▶ Manifold MCMC methods make use of the natural geometry of the underlying statistical model to achieve efficient sampling.
- ▶ Key quantities in information geometry are the Fisher Information (FI) and the connection. Consider a statistical model $S = \{p(\mathbf{y}|\psi) | \psi \in \Psi\}$ for the observed variables \mathbf{y} with parameters ψ . Under quite general conditions, S can be considered a C^∞ manifold, and is called statistical manifold. Let $\mathcal{L} = \log[p(\mathbf{y}|\psi)]$; the FI matrix G of S at ψ is defined as:

$$G(\psi) = E_{p(\mathbf{y}|\psi)} \left[(\nabla_\psi \mathcal{L}) (\nabla_\psi \mathcal{L})^T \right]$$

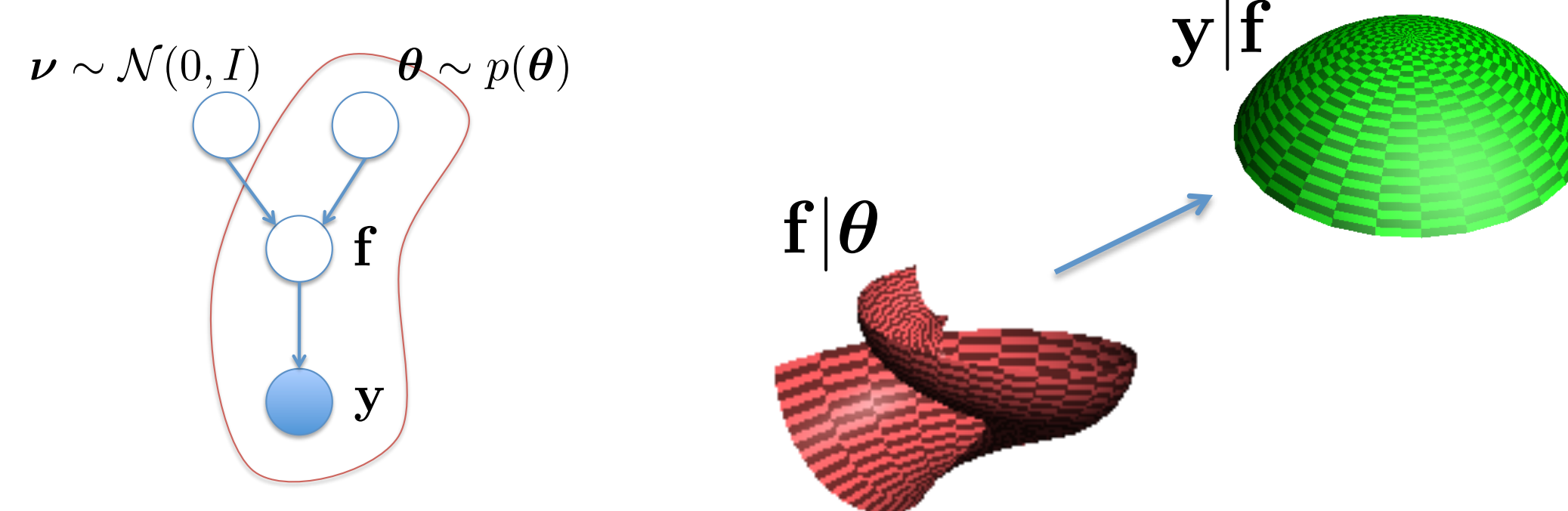
and is the natural metric on S . Christoffel symbols characterize connections on curved manifolds:

$$\Gamma_{kl}^i = \frac{1}{2} \sum_m \left(\frac{\partial g_{mk}}{\partial \psi_l} + \frac{\partial g_{ml}}{\partial \psi_k} - \frac{\partial g_{kl}}{\partial \psi_m} \right)$$

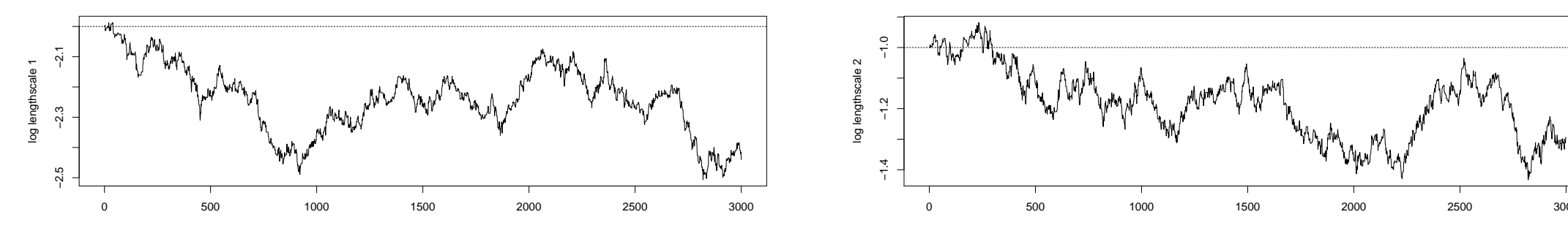
Centered vs Non-centered parameterization

Centered Parameterization

- ▶ Consider the sampling of $p(\mathbf{f}, \theta|\mathbf{y})$. Efficiently sampling from this posterior distribution is complex because of the strong coupling between \mathbf{f} and θ .

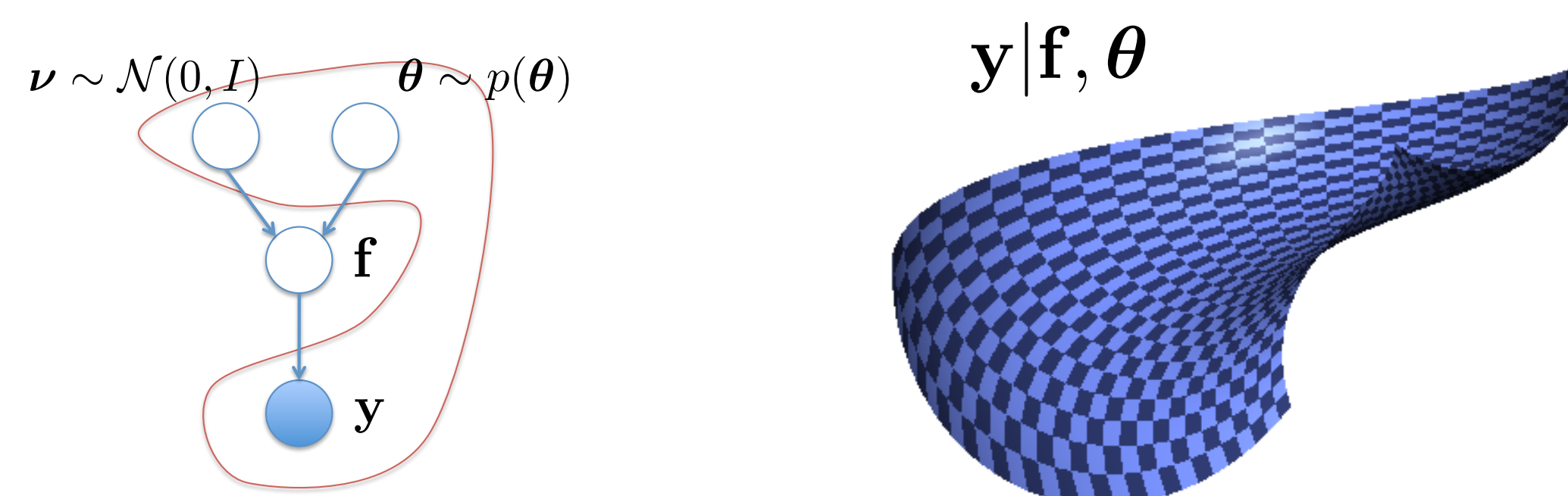


Poor mixing

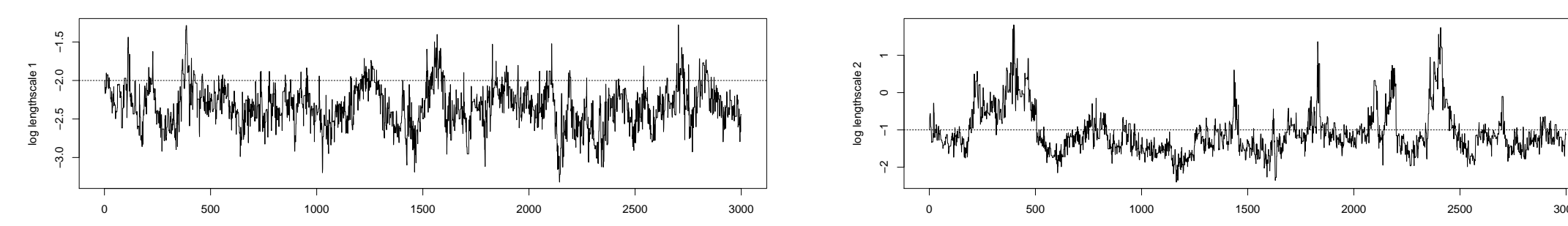


Non-Centered Parameterization

- ▶ If we focus on $p(\nu, \theta|\mathbf{y})$, instead, we can remove the prior correlation.



Better mixing!



Manifold methods for inference in LGMs

- ▶ We consider sampling from the posterior distribution of $p(\nu, \theta|\mathbf{y})$. The expression of the log-joint density reads:

$$\mathcal{L} = \log[h(\mathbf{y})] + \log[g(\zeta)] + \zeta^T \mathbf{u}(\mathbf{y}) \quad \zeta = \zeta(L\nu) = \zeta(\mathbf{f})$$

- ▶ Define

$$R = E_{\mathbf{y}}[\rho\rho^T] \quad \rho = \rho(\theta, \nu, \mathbf{y}) = \nabla_{\mathbf{f}} \zeta(\mathbf{f}) \left(\frac{\nabla_{\zeta} g(\zeta)}{g(\zeta)} + \mathbf{u}(\mathbf{y}) \right)$$

- ▶ The FI results in

$$G_{\nu, \nu} = L^T R L + I \quad G_{\nu, \theta_i} = L^T R \frac{\partial L}{\partial \theta_i} \nu$$

$$G_{\theta_i, \theta_j} = \nu^T \frac{\partial L^T}{\partial \theta_i} R \frac{\partial L}{\partial \theta_j} \nu - \frac{\partial^2 \log[p(\theta)]}{\partial \theta_i \partial \theta_j} \quad G = \begin{pmatrix} G_{\nu, \nu} & G_{\nu, \theta} \\ G_{\theta, \nu} & G_{\theta, \theta} \end{pmatrix}$$

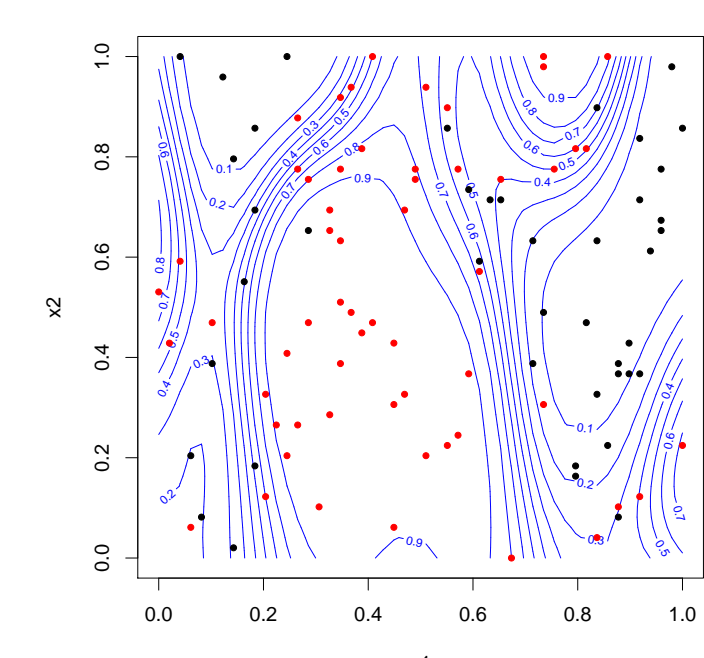
Conclusions and Future work

- ▶ non-centered reparameterization improves sampling efficiency; however, it provides a method to decouple prior correlations, not posterior correlations
- ▶ results for few data show that manifold methods applied to the non-centered parameterization do not improve mixing; preliminary results on larger data sets show that manifold methods improve mixing
- ▶ can we use information geometry to decouple posterior correlations?

References

- [1] O. F. Christensen, G. O. Roberts, and M. Skold. Robust Markov Chain Monte Carlo Methods for Spatial Generalized Linear Mixed Models. *Journal of Computational & Graphical Statistics*, 15(1):1-17, March 2006.
- [2] M. Girolami and B. Calderhead. Riemann Manifold Langevin and Hamiltonian Monte Carlo. *Journal Of The Royal Statistical Society Series B* (with discussion), to appear.
- [3] I. Murray and R. P. Adams. Slice sampling covariance hyperparameters of latent Gaussian models. In *Advances in Neural Information Processing Systems 23 (NIPS 2010)*. 2010.

Results - Logistic regression with Gaussian Process prior



- ▶ bivariate logistic regression with Gaussian Process prior
- ▶ results are for $n = 100$ data points
- ▶ effective sample size (ESS) is computed on 10000 samples from the posterior obtained after 1000 burn-in samples

- ▶ HMC achieves the best ESS but is the most expensive ($\nabla \mathcal{L}$ is expensive)

